



Evaluation of Parametric Statistical Models for Wind Speed Probability Density Estimation

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Outline

- Challenges
- Gaussian Mixture Model
- Expectation-Maximization
- Results and Analysis
- Summary and Conclusions

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Challenges in Intermittent Renewables

- Intermittent generation
- Frequency control
- System inertia
- Power quality



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Statistical Parametric Models

- Parametric Families of Probability Distributions (PFPDs):
 - Wind Speed
 - Rayleigh and Weibull distributions.
- How are they used?
 - Commonly used in power system and microgrid planning.



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Limitations in Common Parametric Distribution Models

- Selecting the appropriate parametric distribution.
- Bimodal or multimodal probability density.
- May be appropriate for a site but may not be appropriate for another.
- Common parametric models are too restrictive and may very well end up in a model mis-specification, which in turn can lead to catastrophic estimates.

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Gaussian Mixture Model (1/2)

- Widely used in:
 - Clustering, machine learning, pattern recognition.
 - Modeling non-Gaussian noise processes, and statistical modeling.
- Gaussian mixtures are often used due to the fact their individual densities are efficiently characterized by first two moments.
- Attributed to have "Universal-approximation" property as proven by Weiner's approximation theorem.
- We are given a dataset $x = \{x_1, ..., x_N\}$, which we believe contains *C* classes. Then, express PDF as as finite convex sum of Gaussian densities.

$$f_X(x_j | \theta) = \sum_{i=1}^{C} \omega_i \frac{1}{\sqrt{2\pi\sigma_i^2}} e^{(-\frac{(x-\mu_i)^2}{2\sigma_i^2})}, \qquad j = \{1, \dots, n\}$$

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Gaussian Mixture Model (2/2)

- Goal:
 - To determine parameters ($\theta = \{\omega_i, \mu_i, \sigma_i^2\}, \forall i = \{1, ..., C\}$), such that likelihood function, $\Pr(\mathbf{x}|\theta, C)$, is maximized.
- Cannot be solved directly:
 - Use Expectation-Maximization (EM) algorithm to maximize the log-likelihood function instead.

• Expectation-Maximization (EM)

- Coined by Dempster et al. in 1977.
- Iterative Solution to Maximum Likelihood Estimation.
- · Used for Mixture Modeling or Clustering.

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Expectation-Maximization (1/2)

• Pros:

- Automated (Unsupervised Algorithm).
- Can be Semi-supervised.
- Guaranteed not to get worse, as it iterates by.
- Cons:
 - Takes time to converge.
 - · Susceptible to over fitting.

Over-fitting

- Very large C maximizes log-likelihood but causes over-fitting.
- Use BIC approach to penalize log-likelihood as *C* increases

$$BIC(C) = -2 \ln \Pr(\mathbf{x}|\hat{\theta}, C) + C \ln(n)$$

where $\hat{\theta} = \{\widehat{\omega_i}, \widehat{\mu_i}, \widehat{\sigma_i^2}\}, i = \{1, ..., C\}$ are the parameter such that log-likelihood is maximized.

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Expectation-Maximization (2/2)

- How to Run Algorithm:
 - Choose Initial values for μ_k , σ_k , ω_k .
 - Calculate responsibilities $\gamma(z_k)$ (closed-form equation exists)
 - Using $\gamma(z_k)$, update values of μ_k , σ_k , π_k (closed-form equation)
 - Repeat until Log-likelihood converges.
- BIC-assisted EM:
 - Do above procedure by varying *C* and running EM until BIC measure gets worse.

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Expectation-Maximization*

- Suppose we are given a dataset $x = \{x_1, ..., x_N\}$, which we believe contains *C* classes, each of which is independent and follow some exponential distribution.
 - Observed Variable:
 - Data, namely $x = \{x_1, ..., x_N\}$
 - Latent (Unobserved) Variable:
 - Which Data belongs to what? i.e.
 - $\Pr(z_k) = \pi_k$.
 - Joint Data, $Y = (\mathbf{x}, Z)$.



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Expectation-Maximization*

- Consists of Two Main Steps:
 - Decide an initial C (number of components)
 - Expectation Step: Compute Posterior Probability
 - Given observed Data x, calculate how likely it is that the complete data is y, i.e. to calculate $Pr(y|D, \theta^m)$

$$\gamma(z_k) = \Pr(z_k = k | \mathbf{x}) \frac{\pi_k N(\mathbf{x} | \mu_k, \sigma_k^2)}{\sum_{l=1}^C \pi_l N(\mathbf{x} | \mu_l, \sigma_l^2)}, k = \{1, ..., C\}$$

- Maximization Step:
 - Given the guess of θ^m , we want to maximize $E\{\log \Pr(x|\theta)\}$.

$$\vec{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \vec{x}_n, \qquad \vec{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\vec{x}_n - \vec{\mu}_k) (\vec{x}_n - \vec{\mu}_k)^{\mathrm{T}},$$

- Termination:
 - Keep iterating until Log-likelihood does not improve significantly.
- BIC-assisted EM:
 - Do above procedure by varying *C* and running EM until BIC measure gets worse.

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Wind Speed Data

- Performance of the proposed GMM for estimating wind speed probability density is assessed via comparisons with conventional parametric probability density models.
- Hourly wind speed data from <u>www.renewables.ninja</u>

Site	Location	Geographical Coordinates (°)		Wind Speed (m/s)				
		Latitude	Longitude	Maximum	Minimum	Mean	Median	Standard Deviation
Site 1	Copenhagen, Denmark	55.69	12.57	23.69	2.41	9.10	9.09	2.82
Site 2	Kiel, Germany	54.46	10.20	19.17	2.63	8.87	8.90	2.73
Site 3	Black Law, Scotland	55.83	-3.82	19.10	1.85	7.40	7.01	2.79
Site 4	Whitelee, Scotland	55.71	-4.34	22.20	2.06	8.43	7.95	3.23
Site 5	Schaw, Scotland	55.46	-4.46	23.26	2.40	9.06	8.63	3.31
Site 6	Malin Head, Ireland	55.38	-7.40	24.32	2.37	10.05	9.53	3.80

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Probability density plots and histogram of wind speed at Denmark, Germany and Ireland



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Probability density plots and histogram of wind speed at Three Sites in Scotland



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Parametric Probability Density Functions

Distribution	Probability Density Function	Domain	Parameters
Rayleigh	$f_X(x) = \frac{x}{b^2} e^{-(x^2/2b^2)}$	$x\in [0,\infty)$	Scale: $b \in (0, \infty)$
Weibull	$f_X(x) = \frac{k}{\lambda} \left(\frac{x}{\lambda}\right)^{k-1} e^{-(x/\lambda)^k}$	$x\in [0,\infty)$	Scale: $\lambda \in (0, \infty)$ Shape: $k \in (0, \infty)$
Gaussian	$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$x \in (-\infty, \infty)$	Mean: $\mu \in (-\infty, \infty)$ Variance: $\sigma^2 \in (0, \infty)$

• Estimated parameters completely define the distribution function for all values of *x*.

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Parameter Estimation of PFPDs

• Estimated using Maximum Likelihood Estimation.

Site	Rayleigh	Weibull		Gaussian	
Sile	b	λ	k	μ	σ^2
Site 1	6.73	10.09	3.50	9.09	7.89
Site 2	6.54	9.81	3.55	8.84	7.35
Site 3	5.62	8.35	2.83	7.43	7.86
Site 4	6.41	9.52	2.78	8.46	10.56
Site 5	6.85	10.21	2.91	9.10	11.09
Site 6	7.64	11.35	2.82	10.10	14.59

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Quantitative Assessment

- Performance of the GMM is assessed against the three selected parametric distributions using:
 - Kolmogorov-Smirnov goodness-of-fit test.
 - Root Mean Square Error.
 - Mean Absolute Error.
- Quantitative way to assess the quality of a given statistical model.

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Kolmogorov-Smirnov Goodness-of-Fit Test

• Used to check whether wind speed data come from a particular statistical model.

Can we *reject* the null hypothesis that the observed data are independently sampled from the six distribution models?

• p-value

- A measure of the evidence against the null hypothesis, with low p-values corresponding to stronger evidence against the null hypothesis.
- *p* < 0.01.

[•] M. DeGroot and M. Schervish, Probability and Statistics. London: Pearson Education, Inc., 4th ed., 2012.

[•] E. L. Lehmann and J. P. Romano, esting Statistical Hypotheses. New York: Springer, 3th ed., 2005.

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p-values resulting from K-S Test

Site	Rayleigh	Weibull	Gaussian	GMM
Site 1	6.41×10^{-80}	24.53×10^{-02}	58.67×10 ⁻⁰²	6.70×10^{-02}
Site 2	2.00×10^{-90}	5.31×10^{-02}	27.27×10^{-02}	5.57×10^{-02}
Site 3	4.18×10^{-39}	6.85×10^{-10}	2.37×10^{-13}	84.42×10^{-02}
Site 4	9.28×10^{-38}	2.59×10^{-09}	5.15×10^{-15}	49.65×10^{-02}
Site 5	9.01×10^{-49}	2.66×10^{-12}	9.36×10^{-14}	18.71×10 ⁻⁰²
Site 6	8.88×10^{-34}	2.76×10^{-08}	7.91×10^{-13}	23.18×10^{-02}

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Error Measures

Root Mean Square Error

Mean Absolute Error

$$RMSE = \sqrt{\frac{1}{t} \sum_{i=1}^{t} (y_i - \hat{y}_i)^2}$$

$$MAE = \frac{1}{t} \sum_{i=1}^{t} |y_i - \hat{y}_i|$$

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Root Mean Square Error

• The proposed method gives the **lowest** RMSE value among all evaluated methods for all six sites.

Site	Rayleigh	Weibull	Gaussian	GMM
Site 1	1.24×10 ⁻² (-)	3.07×10^{-3} (75.2%)	2.95×10^{-3} (76.3%)	2.83×10^{-3} (77.2%)
Site 2	1.19×10^{-2} (-)	4.20×10^{-3} (64.6%)	4.00×10^{-3} (66.3%)	3.81×10 ⁻³ (67.9%)
Site 3	8.85×10 ⁻³ (-)	5.76×10^{-3} (35.0%)	6.45×10^{-3} (27.1%)	3.29×10 ⁻³ (62.8%)
Site 4	8.64×10 ⁻³ (-)	5.88×10^{-3} (32.0%)	6.75×10^{-3} (21.9%)	$4.19 \times 10^{-3} (51.5\%)$
Site 5	9.05×10 ⁻³ (-)	5.77×10^{-3} (36.2%)	6.34×10^{-3} (29.9%)	3.89×10 ⁻³ (56.9%)
Site 6	8.22×10 ⁻³ (-)	5.33×10^{-3} (35.2%)	6.09×10^{-3} (25.9%)	$3.98 \times 10^{-3} (51.6\%)$

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Mean Absolute Error

• The proposed method gives the **lowest** MAE value among all evaluated methods for all six sites.

Site	Rayleigh	Weibull	Gaussian	GMM
Site 1	9.02×10 ⁻³ (-)	2.26×10^{-3} (75.0%)	2.05×10^{-3} (77.3%)	$2.00 \times 10^{-3} (77.9\%)$
Site 2	8.78×10 ⁻³ (-)	2.91×10 ⁻³ (66.8%)	2.79×10^{-3} (68.2%)	2.42×10 ⁻³ (72.4%)
Site 3	5.81×10 ⁻³ (-)	3.94×10^{-3} (32.2%)	4.69×10 ⁻³ (19.3%)	2.34×10 ⁻³ (59.7%)
Site 4	5.82×10 ⁻³ (-)	4.01×10^{-3} (31.0%)	4.85×10^{-3} (16.6%)	$2.66 \times 10^{-3} (54.2\%)$
Site 5	6.57×10 ⁻³ (-)	4.07×10^{-3} (38.1%)	4.51×10^{-3} (31.3%)	2.86×10^{-3} (56.5%)
Site 6	5.56×10 ⁻³ (-)	3.83×10^{-3} (31.2%)	4.43×10^{-3} (20.3%)	2.93×10 ⁻³ (47.3 %)

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Summary and Conclusions

- Common Parametric Distribution Models
- Proposed Gaussian Mixture Model for KDE bandwidth selection
 - BIC-assisted Expectation Minimization Framework
- Probability density plots and histogram of wind speed
- Results of the Kolmogorov-Smirnov Goodness-of-Fit test
- The proposed method gives the lowest RMSE and MAE values among all evaluated methods for all six sites.



Thank You

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