

Performance Analysis of Energy Detection over Mixture Gamma based Fading Channels with Diversity Reception

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Introductory Background | Fading Channels

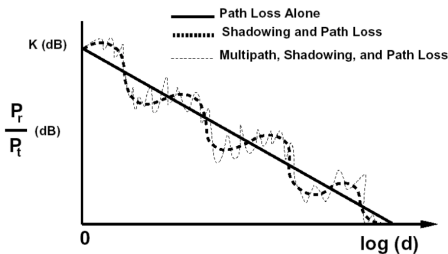
This paper is concerned with providing a **unified approach** to analyzing the performance of **cognitive radio** networks.

- Large-scale (Shadowing) Effect
- Lognormal, Inverse-Gaussian, Gamma.

Introductory Background | Fading Channels

This paper is concerned with providing a **unified approach** to analyzing the performance of **cognitive radio** networks.

- Large-scale (Shadowing) Effect
- Lognormal, Inverse-Gaussian, Gamma.
- Small-scale (Multipath) Effect
- **Conventional:** Rayleigh, Nakagami- m , Weibull- m , Rician.
- **Generalized:** $\alpha - \mu$, $\kappa - \mu$, $\eta - \mu$.



Introductory Background | Composite Fading Channels (1)

Performance analysis over multipath or shadowing scenarios alone is relatively tractable. However, extending the analysis to **composite fading** channels is rather **cumbersome** and **intractable**.

- Examples:

- Nakagami- m /Lognormal (NL) model (**no closed-form**)

$$f_{\gamma}(x) = \frac{2\lambda m^m}{\Gamma(m)\sqrt{2\pi\zeta}} \int_0^{\infty} \frac{x^{m-1}}{\bar{\gamma}^m \sigma^{m+1}} e^{-\frac{mx}{\bar{\gamma}\sigma}} e^{-\frac{(10 \log \sigma)^2}{2\zeta^2}} d\sigma. \quad (1)$$

- $\kappa - \mu$ Shadowed model [1] (**complicated**)

$$f_{\gamma}(\gamma) = \frac{\mu^{\mu} m^m (1 + \kappa)^{\mu}}{\Gamma(\mu) \bar{\gamma} (\mu \kappa + m)^m} \left(\frac{\gamma}{\bar{\gamma}}\right)^{\mu-1} \times \exp\left(-\frac{\mu(1 + \kappa)\gamma}{\bar{\gamma}}\right) {}_1F_1\left(m, \mu; \frac{\mu^2 \kappa (1 + \kappa) \gamma}{\mu \kappa + m} \frac{\gamma}{\bar{\gamma}}\right), \quad (2)$$

[1] J. F. Paris, "Statistical Characterization of κ - μ Shadowed Fading," IEEE Trans. Veh. Technol., vol. 63, no. 2, pp. 518–526, Feb. 2014.

Introductory Background | Composite Fading Channels (2)

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 - The K - and K_G - distributions:
 - Replaced the Lognormal with Gamma and integrate.
 - The RIGD and G -distribution:
 - Replaced the Lognormal with Inverse-Gaussian (IG) and integrate.
- Still **complex** and **not** a general solution.

Introductory Background | Composite Fading Channels (3)

- Another alternative proposed by Atapattu *et al.* [2]
 - Several models were approximated using the mixture Gamma (MG) distribution via gauss-quadrature approximations.
 - Approximates: NL, K , K_G , $\eta - \mu$, $\kappa - \mu$, Hoyt, and Nakagami- m .
- The pdf is given by

$$f_{\gamma}(\gamma) = \sum_{i=1}^K \frac{\alpha_i}{\bar{\gamma}} \frac{x^{\beta_i-1}}{\bar{\gamma}} \exp\left(-\frac{\zeta_i x}{\bar{\gamma}}\right)$$

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- **Pros:** Very tractable, arbitrarily accurate.
- **Cons:** Still not generalizable to all fading models.

[2] S. Atapattu, C. Tellambura, and H. Jiang, "A Mixture Gamma Distribution to Model the SNR of Wireless Channels," *IEEE Trans. Wireless Commun.*, vol. 10, no. 12, pp. 4193–4203, Dec. 2011.

Introductory Background | Energy Detection Performance

- Cognitive radios came as a viable solution to mitigate the spectrum scarcity.
- Spectrum sensing techniques:
 - Energy Detector (ED).
 - Others: Matched Filter, cyclostationarity or feature detection .
- Here we are concerned with the performance of the ED in **generalized and composite** fading channels.

Introductory Background | Energy Detection Performance

- Approaches in literature:
 - A **semi-analytic** solution for the NL was proposed.
 - The K - and K_G - distribution is widely utilized to study the ED performance over RL channels; yet **analytically** cumbersome.
 - The Mixture of Gaussian (MoG) distribution [3] was utilized to study the ED performance.
- Likewise, here we utilize the **mixture distributions** approach, i.e.
 - We utilize the MG distribution to study the ED performance.

[3] O. Alhussein, O. Selim, T. Assaf, S. Muhaidat, J. Liang, and G. K. Karagiannidis, "A Generalized Mixture of Gaussians Model for Fading Channels," in *IEEE Conf. Vehicul. Tech., Glasgow, 2015*.

Energy Detection Performance

- Goal is to study the detection and false-alarm probabilities in MG-based generalized and composite fading channels.
- We consider **square-law combining** (SLC) and **square-law selection** (SLS) diversity schemes.
- Typical configuration:

$$\begin{aligned}\mathcal{H}_0 : y_j(t) &= h_j(t) \\ \mathcal{H}_1 : y_j(t) &= h_j(t) + v_j(t),\end{aligned}\quad (3)$$

- This conditional detection and false-alarm probabilities:

$$P_d = Q_u(\sqrt{2\gamma_j}, \sqrt{\lambda_n}), \quad (4)$$

$$P_f = \frac{\Gamma(u, \frac{\lambda_n}{2})}{\Gamma(u)}, \quad (5)$$

where u is the time-bandwidth product, λ_n is the ED threshold.

Single-antenna Scenario (1)

- Our generalized fading model is written as

$$f_{\gamma}(x) = \sum_{k=1}^K \frac{\alpha_k}{\gamma_0} \left(\frac{x}{\gamma_0}\right)^{\beta_k-1} \exp\left(-\frac{\zeta_k x}{\gamma_0}\right). \quad (6)$$

- Consequently, the detection probability is computed as

$$\bar{P}_{d, MG} = \sum_{k=1}^K \frac{\alpha_k}{\gamma_0} \int_0^{\infty} Q_u(\sqrt{2x}, \sqrt{\lambda_n}) \cdot \left(\frac{x}{\gamma_0}\right)^{\beta_k-1} e^{-\frac{\zeta_k x}{\gamma_0}} dx. \quad (7)$$

Single-antenna Scenario (2)

- Using Theorem 1 in [4], we obtain

$$\begin{aligned} \bar{P}_{d,MG} = & \sum_{k=1}^K \frac{\alpha_k \Gamma(\beta_k) \Gamma(u, \frac{\lambda_n}{2})}{\Gamma(u) \zeta_k^{\beta_k}} + \sum_{k=1}^K \sum_{l=0}^{\beta_k-1} \frac{\alpha_k \Gamma(\beta_k)}{\gamma_0^{\beta_k}} \\ & \left(\frac{\lambda_n}{2}\right)^u {}_1F_1(l+1, u+1, \frac{\lambda_n/2}{1+1\frac{\zeta_k}{\gamma_0}}) \\ & \times \frac{1}{u! \left(\frac{\zeta_k}{\gamma_0}\right)^{\beta_k-l} \left(1 + \frac{\zeta_k}{\gamma_0}\right)^{l+1} \exp\left(\frac{\lambda_n}{2}\right)}, \end{aligned} \quad (8)$$

[4] P. Sofotasios, M. Valkama, T. Tsiftsis, Y. Brychkov, S. Freear, and G. Karagiannidis, "Analytic solutions to a Marcum Q-function-based integral and application in energy detection of unknown signals over multipath fading channels," in IEEE Cogn. Radio Oriented Wireless Netw. and Commun. (CROWNCOM), 2014, pp. 260–265.

Square-Law Combining (1)

- Under SLC, the received signals from each branch are integrated, squared, and then summed up.

$$\gamma_{\Sigma} = \sum_{l=1}^L \gamma_l. \quad (9)$$

- We need to obtain pdf of γ_{Σ} in order to derive $P_{d,\Sigma}$.

$$f_{\gamma_{\Sigma}}^{(2)} = \int_0^{\gamma} f_{\gamma_1}(x) f(\gamma - x) dx = \sum_{i=1}^K \sum_{j=1}^K \frac{\alpha_i \alpha_j}{\gamma_0^{\beta_i + \beta_j}} \times \int_0^{\gamma} x^{\beta_i - 1} e^{-\frac{\zeta_i}{\gamma_0} x} (\gamma - x)^{\beta_j - 1} e^{-\frac{\zeta_j}{\gamma_0} (\gamma - x)} dx. \quad (10)$$

Square-Law Combining (2)

- Under SLC, the received signals from each branch are integrated, squared, and then summed up.

$$\gamma_{\Sigma} = \sum_{l=1}^L \gamma_l. \quad (11)$$

- The pdf of γ_{Σ} was derived

$$f_{\gamma_{\Sigma}}^{(2)} |_{(\zeta_i = \zeta_j)} = \sum_{i=1}^K \sum_{j=1}^K \frac{\alpha_i \alpha_j}{\gamma_0^{\beta_i + \beta_j}} \frac{\Gamma(\beta_i) \Gamma(\beta_j)}{\Gamma(\beta_i + \beta_j)} e^{-\frac{\zeta_j}{\gamma_0} \gamma} \gamma^{\beta_i + \beta_j - 1}. \quad (12)$$

Square-Law Combining (3)

$$\begin{aligned}
 f_{\gamma\Sigma}^{(2)}|_{(\zeta_i \neq \zeta_j)} &= \sum_{i=1}^K \sum_{j=1}^K \sum_{l=0}^{\beta_j-1} \binom{\beta_j-1}{l} \frac{(-1)^l \alpha_i \alpha_j}{\gamma_0^{\beta_j-l} (\zeta_i - \zeta_j)^{\beta_i+l}} \\
 &\times \gamma^{\beta_j-l-1} e^{-\frac{\zeta_j}{\gamma_0} \gamma} \gamma \left(\beta_i + l, \frac{\gamma(\zeta_i - \zeta_j)}{\gamma_0} \right), \quad (13)
 \end{aligned}$$

while, $f_{\gamma\Sigma}^{(2)} = f_{\gamma\Sigma}^{(2)}|_{(\zeta_i = \zeta_j)} + f_{\gamma\Sigma}^{(2)}|_{(\zeta_i \neq \zeta_j)}$.

- Note: This method works for higher diversity orders.

Square-Law Combining (4)

- Thus \bar{P}_{d,Σ_2} was derived as

$$\begin{aligned} \bar{P}_{d,\Sigma}^{(2)} | (\zeta_i = \zeta_j) &= \sum_{i=1}^K \sum_{j=1}^K \alpha_i \alpha_j \Gamma(\beta_i) \Gamma(\beta_j) \left[\frac{\Gamma(u, \frac{\lambda}{2})}{\zeta_j^{\beta_i + \beta_j} \Gamma(u)} \right. \\ &+ \left. \sum_{n=0}^{\beta_i + \beta_j - 1} \frac{\gamma_0^{-n} (\frac{\lambda}{2})^u {}_1F_1(n+1, u+1, \frac{\frac{\lambda}{2}}{1 + \frac{\zeta_j}{\gamma_0}})}{u! (\zeta_j)^{\beta_i + \beta_j - n} (1 + \frac{\zeta_j}{\gamma_0})^{n+1} \exp(\frac{\lambda}{2})} \right] (14) \end{aligned}$$

Square-Law Combining (5)

$$\begin{aligned}
 \bar{P}_{d,\Sigma}^{(2)}|_{(\zeta_i \neq \zeta_j)} &= \sum_{i=1}^K \sum_{j=1}^K \sum_{l=0}^{\beta_j-1} \binom{\beta_j-1}{l} \frac{(-1)^l \alpha_i \alpha_j \Gamma(\beta_i + l)}{\gamma_0^{\beta_j-l} (\zeta_i - \zeta_j)^{\beta_i+l}} \\
 &\times \left[\frac{\Gamma(\beta_j - l) \Gamma(u, \frac{\lambda}{2})}{(\frac{\zeta_i}{\gamma_0})^{\beta_j-l} \Gamma(u)} + \sum_{n=0}^{\beta_j-l-1} \frac{(\frac{\lambda}{2})^u \Gamma(\beta_j - l)}{u! (\frac{\zeta_i}{\gamma_0})^{\beta_j-l-n}} \frac{{}_1F_1(n+1, u+1, \frac{\frac{\lambda}{2}}{1+\frac{\zeta_i}{\gamma_0}})}{(1+\frac{\zeta_i}{\gamma_0})^{n+1} \exp(\frac{\lambda}{2})} \right] \\
 &- \sum_{i=1}^K \sum_{j=1}^K \sum_{l=0}^{\beta_j-1} \sum_{t=0}^{\beta_i+l-1} \binom{\beta_j-1}{l} \frac{(-1)^l \alpha_i \alpha_j \Gamma(\beta_i + l)}{t! \gamma_0^{\beta_j-l+t} (\zeta_i - \zeta_j)^{\beta_i+l-t}} \\
 &\times \left[\frac{\Gamma(\beta_j + t - l) \Gamma(u, \frac{\lambda}{2})}{(\frac{\zeta_i}{\gamma_0})^{\beta_j+t-l} \Gamma(u)} + \sum_{n=0}^{\beta_j+t-l-1} \frac{(\frac{\lambda}{2})^u \Gamma(\beta_j + t - l) {}_1F_1(n+1, u+1, \frac{\frac{\lambda}{2}}{1+\frac{\zeta_i}{\gamma_0}})}{u! (\frac{\zeta_i}{\gamma_0})^{\beta_j+t-l-n} (1+\frac{\zeta_i}{\gamma_0})^{n+1} \exp(\frac{\lambda}{2})} \right].
 \end{aligned} \tag{15}$$

Square-Law Selection (1)

- Under SLS, the branch with the maximum γ_j is selected as follows

$$\gamma_{\text{SLS}} = \max_{j=1,\dots,L} (\gamma_j). \quad (16)$$

- Under \mathcal{H}_0 , the false-alarm probability for the SLS scheme can be expressed as

$$P_{f,\text{SLS}} = 1 - \Pr(\gamma_{\text{SLS}} < \lambda_n | \mathcal{H}_0). \quad (17)$$

- Substituting (16) in (17), we obtain

$$P_{f,\text{SLS}} = 1 - \Pr(\max(\gamma_1, \gamma_2, \dots, \gamma_L) < \lambda_n | \mathcal{H}_0). \quad (18)$$

Square-Law Selection (2)

- Accordingly, this translates to

$$P_{f,SLS} = 1 - [1 - P_f]^L. \quad (19)$$

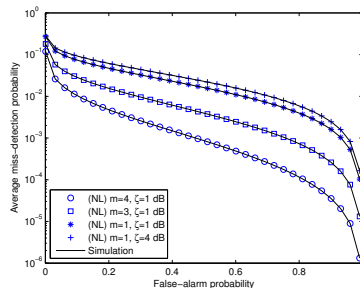
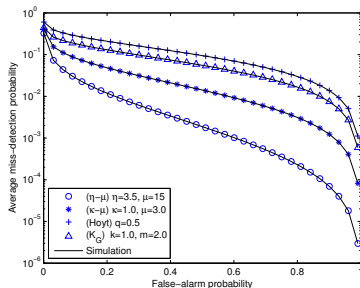
- Similarly, the unconditional probability of detection over the AWGN channel is obtained by

$$P_{d,SLS} = 1 - \prod_{j=1}^L \left[1 - Q_u(\sqrt{2\gamma_j}, \sqrt{\lambda_n}) \right]. \quad (20)$$

- Hence, averaging (20) over (7) yields the unconditional probability of detection under the SLS scheme, $\bar{P}_{d,SLS}$, which is given by

$$\bar{P}_{d,SLS} = 1 - \prod_{j=1}^L [1 - P_{d,MG}]. \quad (21)$$

Simulation Results



(Left) depicts $1 - P_d$ versus P_{fa} (ROC) curves for different fading conditions with no diversity. (Right) depicts the ROC curves over several scenarios of the composite NL fading channel with SLC diversity scheme with $L = 2$.

Conclusions

- We proposed a unified framework for the performance analysis of an energy detector in **generalized and composite** MG-based fading channels.
- Novel analytical expressions for the average detection probability have been derived for both the single-antenna case and the multiple-antenna case with SLC and SLS schemes.
 - Derived expressions are **generalized** in terms of fading characterization and **algebraically versatile**.
- Notes:
 - This paper can be found in the [ArXiv.org](https://arxiv.org/abs/1510.05594) repository ([arXiv:1510.05594](https://arxiv.org/abs/1510.05594)).
 - An extended journal edition of the paper is found as an *early-access* IEEE **TVT** transaction.